Sparse multioutput radial basis function network construction using combined locally regularised orthogonal least square and D-optimality experimental design

S. Chen, X. Hong and C.J. Harris

Abstract: A construction algorithm for multioutput radial basis function (RBF) network modelling is introduced by combining a locally regularised orthogonal least squares (LROLS) model selection with a D-optimality experimental design. The proposed algorithm aims to achieve maximised model robustness and sparsity via two effective and complementary approaches. The LROLS method alone is capable of producing a very parsimonious RBF network model with excellent generalisation performance. The D-optimality design criterion enhances the model efficiency and robustness. A further advantage of the combined approach is that the user only needs to specify a weighting for the D-optimality cost in the combined RBF model selecting criterion and the entire model construction procedure becomes automatic. The value of this weighting does not influence the model selection procedure critically and it can be chosen with ease from a wide range of values.

1 Introduction

The radial basis function (RBF) network has widely been studied [1–7]. For single-output nonlinear data modelling or regression, the orthogonal least squares (OLS) algorithm [4, 8] provides an effective means to construct parsimonious RBF networks with good generalisation performance. The parsimonious principle alone, however, is not entirely immune to over-fitting. If data are highly noisy, small models constructed may still fit into noise. A useful technique for overcoming over-fitting is regularisation [9–12]. From the Bayesian learning viewpoint, regularisation is equivalent to adopting a hyperparameter approach [13, 14], and recent work [15, 16] has combined the OLS algorithm with an individually regularised approach to derive an efficient single-output locally regularised OLS (LROLS) algorithm. Optimal experimental designs [17] have been used to construct smooth model response surfaces based on the setting of the experimental variables under well controlled experimental conditions. In optimal design, model adequacy is evaluated by design criteria that are statistical measures of goodness of experimental designs by virtue of design efficiency and experimental effort. For regression models, quantitatively, model adequacy is measured as a function of the eigenvalues of the design matrix. The D-optimality design criterion [17] is most effective in optimising the parameter efficiency and model robustness via the maximisation of the determinant of the design matrix. The traditional nonlinear model structure determination based on optimal experimental designs is inherently inefficient and computationally prohibitive. Recently, effective model construction algorithms has been proposed for single-output nonlinear modelling based on the computationally efficient OLS and LROLS algorithms, coupled with the D-optimality experimental design [18, 19].

For the construction of multioutput RBF networks, one approach is to fit multiple single-output models as, for example, in the work [20], and an alternative is to construct a single multioutput RBF network model as, for example, in the work [21]. The latter approach has an advantage: a selected RBF term must be significant in explaining all the outputs, and this can result in a smaller number of regressors than the former approach to achieve the same modelling accuracy. Recent work [22] has combined the local regularisation approach with the multioutput OLS regression. This paper proposes to combine the multioutput LROLS algorithm [22] with the D-optimality experimental design. Computational efficiency of the resulting algorithm is ensured by the orthogonal forward selection procedure. The local regularisation enforces model sparsity and avoids over-fitting while the D-optimality design optimises model efficiency and parameter robustness. The coupling effects of these two approaches in the combined algorithm further enhance each other. The end result is an efficient yet simple algorithm for constructing sparse multioutput RBF models that generalise well, especially under highly noisy learning conditions. Moreover, the model construction process becomes fully automatic, and there is only one user

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specified quantity which has no critical influence on the model selection procedure.

2 Multioutput radial basis function network

Consider the general discrete-time nonlinear system represented by the nonlinear model [23]

$$\mathbf{y}(k) = \mathbf{f}(\mathbf{y}(k-1), \ldots, \mathbf{y}(k-n_y), \mathbf{u}(k-1), \ldots, \mathbf{u}(k-n_u)) + \mathbf{e}(k) = \mathbf{f}(\mathbf{x}(k)) + \mathbf{e}(k) \quad (1)$$

where

$$\mathbf{u}(k) = [u_1(k) \cdots u_{n_i}(k)]^T \in \mathcal{R}^{n_i} \quad (2)$$

and

$$\mathbf{y}(k) = [y_1(k) \cdots y_{n_o}(k)]^T \in \mathcal{R}^{n_o} \quad (3)$$

are the system input and output vector variables with dimensions n_i and n_o, respectively, n_u, and n_y, are positive integers representing the lags in $\mathbf{u}(k)$ and $\mathbf{y}(k)$, respectively,

$$\mathbf{e}(k) = [e_1(k) \cdots e_{n_o}(k)]^T \in \mathcal{R}^{n_o} \quad (4)$$

is the system white noise vector with covariance $\text{Cov}[\mathbf{e}(k)] = \sigma_e^2 \mathbf{I}_{n_o}$ and $\mathbf{I}_{n_o}$ being the $n_o \times n_o$ identity matrix,

$$\mathbf{x}(k) = [\mathbf{y}^T(k-1) \cdots \mathbf{y}^T(k-n_y) \mathbf{u}^T(k-1) \cdots \mathbf{u}^T(k-n_u)]^T \quad (5)$$

denotes the system "input" vector; and $\mathbf{f}(\bullet)$ is the unknown n_o-dimensional system mapping. The system model (1) is to be identified from an N-sample observation data set $\{\mathbf{x}(k), \mathbf{y}(k)\}_{k=1}^N$ using some suitable functional which can approximate $\mathbf{f}(\bullet)$ with arbitrary accuracy. One class of such functionals is the RBF network model of the form

$$y_i(k) = \hat{y}_i(k) + e_i(k) = \sum_{j=1}^{M} \theta_{j,i} \phi_j(\mathbf{x}(k)) + e_i(k), \quad 1 \le k \le N \quad (6)$$

for $1 \le i \le n_o$, where $e_i(k)$ is the error between $y_i(k)$ and the ith model output $\hat{y}_i(k)$, $\theta_{j,i}$ are the RBF weights, the RBF kernels or regressors

$$\phi_j(\mathbf{x}(k)) = \phi(\|\mathbf{x}(k) - \mathbf{c}_j\|; \rho_j) \quad (7)$$

$\mathbf{c}_j$ are the RBF centers and ρ_j the positive width parameters. Typically, each training data $\mathbf{x}(k)$ is considered as a candidate RBF centre, and the total number of candidate regressors in this case is $M = N$. Typical choices of nonlinearity $\phi(\bullet)$ are

$$\begin{aligned} \phi(v) &= v^2 \log(v), && \text{thin-plate-spline} \\ \phi(v; \rho) &= \exp\left(-\frac{v^2}{2\rho^2}\right), && \text{gaussian} \\ \phi(v; \rho) &= (v^2 + \rho^2)^{1/2}, && \text{multiquadric} \\ \phi(v; \rho) &= \frac{1}{\sqrt{v^2 + \rho^2}}, && \text{inverse multiquadric} \end{aligned} \quad (8)$$

The multioutput RBF network model (6) can be written in a more concise form as

$$\mathbf{y}_i = \mathbf{\Phi} \boldsymbol{\theta}_i + \mathbf{e}_i, \quad 1 \le i \le n_o \quad (9)$$

by defining

$$\mathbf{y}_i = \begin{bmatrix} y_i(1) \\ y_i(2) \\ \vdots \\ y_i(N) \end{bmatrix}, \quad \mathbf{e}_i = \begin{bmatrix} e_i(1) \\ e_i(2) \\ \vdots \\ e_i(N) \end{bmatrix}, \quad \theta_i = \begin{bmatrix} \theta_{1,i} \\ \theta_{2,i} \\ \vdots \\ \theta_{M,i} \end{bmatrix} \quad (10)$$

for $1 \le i \le n_o$, and

$$\mathbf{\Phi} = [\boldsymbol{\phi}_1 \quad \boldsymbol{\phi}_2 \quad \cdots \quad \boldsymbol{\phi}_M] \quad (11)$$

with

$$\phi_j = [\phi_j(\mathbf{x}(1)) \phi_j(\mathbf{x}(2)) \cdots \phi_j(\mathbf{x}(N))]^T, \quad 1 \le j \le M \quad (12)$$

Further, define

$$\mathbf{Y} = [\mathbf{y}_1 \quad \mathbf{y}_2 \quad \cdots \quad \mathbf{y}_{n_o}], \quad \mathbf{\Theta} = [\boldsymbol{\theta}_1 \quad \boldsymbol{\theta}_2 \quad \cdots \quad \boldsymbol{\theta}_{n_o}], \quad \mathbf{E} = [\mathbf{e}_1 \quad \mathbf{e}_2 \quad \cdots \quad \mathbf{e}_{n_o}] \quad (13)$$

The RBF network model (6) is given in the matrix form as

$$\mathbf{Y} = \mathbf{\Phi}\mathbf{\Theta} + \mathbf{E} \quad (14)$$

Let an orthogonal decomposition of the regression matrix $\mathbf{\Phi}$ be

$$\mathbf{\Phi} = \mathbf{W}\mathbf{A} \quad (15)$$

where

$$\mathbf{A} = \begin{bmatrix} 1 & a_{1,2} & \cdots & a_{1,M} \\ 0 & 1 & \ddots & \vdots \\ \vdots & \ddots & \ddots & a_{M-1,M} \\ 0 & \cdots & 0 & 1 \end{bmatrix} \quad (16)$$

and

$$\mathbf{W} = [\mathbf{w}_1 \quad \mathbf{w}_2 \quad \cdots \quad \mathbf{w}_M] \quad (17)$$

which satisfies $\mathbf{w}_j^T \mathbf{w}_l = 0$, if $j \ne l$. The RBF model (14) can alternatively be expressed as

$$\mathbf{Y} = \mathbf{W}\mathbf{G} + \mathbf{E} \quad (18)$$

where the orthogonal weight matrix

$$\mathbf{G} = [\mathbf{g}_1 \quad \mathbf{g}_2 \quad \cdots \quad \mathbf{g}_{n_o}] \quad (19)$$

with

$$\mathbf{g}_i = [g_{1,i} \quad g_{2,i} \quad \cdots \quad g_{M,i}]^T, \quad 1 \le i \le n_o \quad (20)$$

and $\mathbf{G}$ satisfies the triangular system

$$\mathbf{A}\mathbf{\Theta} = \mathbf{G} \quad (21)$$

Knowing $\mathbf{A}$ and $\mathbf{G}$, $\mathbf{\Theta}$ can readily be solved from (21).

3 Multioutput LROLS algorithm with D-optimality design

Before discussing this combined multioutput model construction algorithm, its two components, the LROLS algorithm and the D-optimality experimental design, are briefly discussed.

3.1 LROLS algorithm

The multioutput LROLS algorithm is based on the following regularised error criterion [22]:

$$J_R(\mathbf{G}, \boldsymbol{\lambda}) = \text{trace}\,(\mathbf{E}^T\mathbf{E} + \mathbf{G}^T\mathbf{\Lambda}\mathbf{G}) = \sum_{i=1}^{n_o} (\mathbf{e}_i^T \mathbf{e}_i + \mathbf{g}_i^T \mathbf{\Lambda} \mathbf{g}_i) = \sum_{i=1}^{n_o} \mathbf{e}_i^T \mathbf{e}_i + \sum_{j=1}^{M} \left(\sum_{i=1}^{n_o} g_{j,i}^2 \right) \lambda_j \quad (22)$$

where $\boldsymbol{\lambda} = [\lambda_1 \ \lambda_2 \ \cdots \ \lambda_M]^T$ is the regularisation parameter vector, and the diagonal matrix $\mathbf{\Lambda} = \text{diag}\{\lambda_1, \lambda_2, \cdots, \lambda_M\}$. The original multioutput OLS algorithm [21] can be

viewed as a special case with $\lambda_j = 0$, $\forall j$. After some simplification, the criterion (22) can be expressed as [22]

$$\text{trace}\,(\mathbf{E}^T\mathbf{E} + \mathbf{G}^T\mathbf{\Lambda}\mathbf{G}) = \text{trace}\,(\mathbf{Y}^T\mathbf{Y} - \mathbf{G}^T(\mathbf{W}^T\mathbf{W} + \mathbf{\Lambda})\mathbf{G}) \tag{23}$$

or

$$\text{trace}\,(\mathbf{E}^T\mathbf{E} + \mathbf{G}^T\mathbf{\Lambda}\mathbf{G}) = \sum_{i=1}^{n_o}\mathbf{y}_i^T\mathbf{y}_i - \sum_{j=1}^{M}\left(\sum_{i=1}^{n_o} g_{j,i}^2\right)(\mathbf{w}_j^T\mathbf{w}_j + \lambda_j) \tag{24}$$

Normalising (23) by trace $(\mathbf{Y}^T\mathbf{Y})$ yields

$$\frac{\text{trace}\,(\mathbf{E}^T\mathbf{E} + \mathbf{G}^T\mathbf{\Lambda}\mathbf{G})}{\text{trace}\,(\mathbf{Y}^T\mathbf{Y})} = 1 - \sum_{j=1}^{M}\frac{\left(\sum_{i=1}^{n_o} g_{j,i}^2\right)(\mathbf{w}_j^T\mathbf{w}_j + \lambda_j)}{\text{trace}\,(\mathbf{Y}^T\mathbf{Y})} \tag{25}$$

Define the regularised error reduction ratio due to the regressor $\mathbf{w}_l$ as

$$[\text{rerr}]_l = \frac{\left(\sum_{i=1}^{n_o} g_{l,i}^2\right)(\mathbf{w}_l^T\mathbf{w}_l + \lambda_l)}{\text{trace}\,(\mathbf{Y}^T\mathbf{Y})} \tag{26}$$

Based on this ratio, significant regressors can be selected in a forward-regression procedure [22]. At the lth stage, a regressor is chosen as the lth term of the subset model if it produces the largest $[\text{rerr}]_l$ among the remaining $M - l + 1$ candidates, and the selection is terminated at the M_sth stage when

$$1 - \sum_{l=1}^{M_s}[\text{rerr}]_l < \xi \tag{27}$$

is satisfied, where $0 < \xi < 1$ is a chosen tolerance. This produces a sparse model containing $M_s (\ll M)$ significant regressors. The detailed algorithm selection procedure can be found in [22]. Notice that, in the selection procedure, if $\mathbf{w}_l^T\mathbf{w}_l$ is too small (near zero), this term will not be selected. Thus, any ill-conditioning or singular situations can automatically be avoided. The Bayesian evidence procedure [13] can readily be extended to the multioutput case and thus used to "optimise" the regularisation parameters. This leads to the updating formulas for the regularisation parameters [22]

$$\lambda_j^{\text{new}} = \frac{\gamma_j^{\text{old}}}{N - \gamma^{\text{old}}}\frac{\sum_{i=1}^{n_o}\mathbf{e}_i^T\mathbf{e}_i}{\sum_{i=1}^{n_o} g_{j,i}^2}, \quad 1 \le j \le M \tag{28}$$

where

$$\gamma_j = \frac{\mathbf{w}_j^T\mathbf{w}_j}{\lambda_j + \mathbf{w}_j^T\mathbf{w}_j} \tag{29}$$

and

$$\gamma = \sum_{j=1}^{M}\gamma_j \tag{30}$$

Usually a few iterations (typically 10 to 30) are sufficient to find an optimal $\boldsymbol{\lambda}$.

It is worth emphasising that, for this multioutput LROLS algorithm, the choice of ξ is less critical than the original OLS algorithm. This is because multiple regularisers enforce sparsity. If, for example, ξ is chosen too small, those unnecessarily selected terms will have a very large λ_l associated with each of them, effectively forcing their weights to zero [15, 16]. Nevertheless, an appropriate value for ξ is desired. Alternatively, the Akaike information criterion (AIC) [24, 25] can be adopted to terminate the subset model selection process. The AIC can be viewed as a model structure regularisation by conditioning the model size using a penalty term to penalise large sized models. However, the use of AIC or other information based criteria in forward regression only affects the stopping point of the model selection, but does not penalises the regressor that may cause poor model performance (e.g. too large variance of parameter estimate or ill-posedness of the regression matrix), if it is selected. Or simply the penalty term in AIC does not determine which regressor should be selected. Optimal experimental design criteria offer better solutions as they are directly linked to model efficiency and parameter robustness.

3.2 D-optimality experimental design

In experimental design, the data covariance matrix $\mathbf{\Phi}^T\mathbf{\Phi}$ is called the design matrix. The least squares (LS) estimate of $\mathbf{\Theta}$ is given by $\hat{\mathbf{\Theta}} = (\mathbf{\Phi}^T\mathbf{\Phi})^{-1}\mathbf{\Phi}^T\mathbf{Y}$. Assume that (14) represents the true data generating process and $\mathbf{\Phi}^T\mathbf{\Phi}$ is nonsingular. Then the estimate $\hat{\mathbf{\Theta}}$ is unbiased and the covariance matrix of the estimate is determined by the design matrix

$$\begin{cases} E[\hat{\mathbf{\Theta}}] = \mathbf{\Theta}, \\ \text{Cov}[\hat{\mathbf{\Theta}}] \propto (\mathbf{\Phi}^T\mathbf{\Phi})^{-1} \end{cases} \tag{31}$$

It is well known that models based on the LS estimate tend to be unsatisfactory for an ill-conditioned regression or design matrix. The condition number of the design matrix is given by

$$C = \frac{\max\{\kappa_i, \quad 1 \le i \le M\}}{\min\{\kappa_i, \quad 1 \le i \le M\}} \tag{32}$$

with κ_i, $1 \le i \le M$, being the eigenvalues of $\mathbf{\Phi}^T\mathbf{\Phi}$. Too large a condition number will result in unstable LS parameter estimate while a small condition number improves model robustness. The D-optimality design criterion maximises the determinant of the design matrix for the constructed model. Specifically, let $\mathbf{\Phi}_{M_s}$ be a column subset of $\mathbf{\Phi}$ representing a constructed M_s-term subset model. According to the D-optimality criterion, the selected subset model is the one that maximises $\det(\mathbf{\Phi}_{M_s}^T\mathbf{\Phi}_{M_s})$. This helps to prevent the selection of an oversized ill-posed model and the problem of high parameter estimate variances. Thus, the D-optimality design is aimed to optimise model efficiency and parameter robustness.

The optimal experimental designs 'however' do not provide means of parameter estimates and have to rely on the LS or regularised LS methods for model parameter estimate. It is straightforward to verify that maximising $\det(\mathbf{\Phi}_{M_s}^T\mathbf{\Phi}_{M_s})$ is identical to maximising $\det(\mathbf{W}_{M_s}^T\mathbf{W}_{M_s})$ or, equivalently, minimising $-\log\det(\mathbf{W}_{M_s}^T\mathbf{W}_{M_s})$ [18]. Note that

$$\begin{aligned}\det(\mathbf{\Phi}^T\mathbf{\Phi}) &= \prod_{i=1}^{M}\kappa_i = \det(\mathbf{A}^T)\det(\mathbf{W}^T\mathbf{W})\det(\mathbf{A}) \\ &= \det(\mathbf{W}^T\mathbf{W}) = \prod_{i=1}^{M}\mathbf{w}_i^T\mathbf{w}_i\end{aligned} \tag{33}$$

and

$$-\log\left(\det(\mathbf{W}^T\mathbf{W})\right) = \sum_{i=1}^{M} -\log(\mathbf{w}_i^T\mathbf{w}_i) \tag{34}$$

By utilising the additive property of (34) the D-optimality design criterion can be incorporated naturally and efficiently with the orthogonal forward regression procedure.

3.3 Combined LROLS and D-optimality algorithm

The combined LROLS and D-optimality algorithm can be viewed as based on the combined criterion of

$$J_C(\mathbf{G}, \boldsymbol{\lambda}, \beta) = J_R(\mathbf{G}, \boldsymbol{\lambda}) + \beta \sum_{j=1}^{M} -\log(\mathbf{w}_j^T \mathbf{w}_j) \tag{35}$$

where β is a fixed small positive weighting for the D-optimality cost. In this combined algorithm the updating of the model weights and regularisation parameters is exactly as in the LROLS algorithm, but the selection is according to the combined regularised error reduction ratio defined as

$$[\text{crerr}]_l = \frac{(\sum_{i=1}^{n_o} g_{l,i}^2)(\mathbf{w}_l^T \mathbf{w}_l + \lambda_l) + \beta \log(\mathbf{w}_l^T \mathbf{w}_l)}{\text{trace}\,(\mathbf{Y}^T \mathbf{Y})} \tag{36}$$

and the selection is terminated with an M_s-term model when

$$[\text{crerr}]_l \leq 0 \quad \text{for} \quad M_s + 1 \leq l \leq M \tag{37}$$

The iterative RBF model selection procedure can now be summarised.

Initialisation: Set λ_j, $1 \leq j \leq M$, to the same small positive value (e.g. 0.001) and choose a fixed β. Set iteration index $I = 1$.

Step 1: Given the current $\boldsymbol{\lambda}$, select a subset model with M_I terms using the forward regression based on $[\text{crerr}]_I$.

Step 2: Update $\boldsymbol{\lambda}$ using (28)–(30) with $M = M_I$. If $\boldsymbol{\lambda}$ remains sufficiently unchanged in two successive iterations or a preset maximum iteration number is reached, stop; otherwise set $I = I + 1$ and go to step 1.

The introduction of the D-optimality cost into the algorithm further enhances the efficiency and robustness of the selected subset model and as a consequence the combined algorithm can often produce sparser models with equally good generalisation properties, compared with the LROLS algorithm. Note that the model selection procedure is simplified and it is no longer necessary to specify the tolerance ξ, as the algorithm automatically terminates when condition (37) is reached. Unlike the combined OLS and D-optimality algorithm [18], the value of weighting β does not critically influence the performance of this combined LROLS and D-optimality algorithm and β can be chosen with ease from a large range of values. This will be demonstrated in the following modelling examples. It should also be emphasised that the computational complexity of this algorithm is not significantly more than that of the OLS algorithm. This is simply because after the first iteration, which has a complexity of the OLS algorithm, the model set contains only $M_1 (\ll M)$ terms, and the complexity of the subsequent iteration decreases dramatically. After a few iterations, typically the model set will converge to a constant size of very small M_s. A few more iterations will ensure the convergence of $\boldsymbol{\lambda}$. Thus, this combined LROLS and D-optimality design algorithm offers an efficient procedure to construct sparse multioutput RBF models with excellent generalisation performance without the need to apply costly cross-validation.

4 Nonlinear system modelling examples

Three examples illustrate the effectiveness of the multioutput LROLS algorithm with the D-optimality design and to compare it with the combined OLS algorithm and D-optimality design. The RBF network model used in the simulation employed the thin-plate-spline nonlinearity.

4.1 Simulated two-output time series process

The data set contained 1000 noisy observations which were generated using the model

$$\begin{aligned} y_1(k) &= 0.1 \sin(\pi y_2(k-1)) \\ &\quad + (0.8 - 0.5 \exp(-y_1^2(k-1)))y_1(k-1) \\ &\quad - (0.3 + 0.9 \exp(-y_1^2(k-1)))y_1(k-2) + \epsilon_1(k) \\ y_2(k) &= 0.6 y_2(k-1) + 0.2 y_2(k-1) y_2(k-2) \\ &\quad + 1.2 \tanh(y_1(k-2)) + \epsilon_2(k) \end{aligned} \tag{38}$$

given the initial conditions $y_1(0) = y_1(-1) = y_2(0) = y_2(-1) = 0$, where the zero-mean gaussian noise $\epsilon(k) = [\epsilon_1(k) \epsilon_2(k)]^T$ had a covariance $0.04\mathbf{I}_2$. The first 500 data samples were used for training and the other 500 samples for validating the obtained model. The underlying dynamics of the simulated time series was governed by

$$\begin{aligned} y_{d1}(k) &= 0.1 \sin(\pi y_{d2}(k-1)) \\ &\quad + (0.8 - 0.5 \exp(-y_{d1}^2(k-1)))y_{d1}(k-1) \\ &\quad - (0.3 + 0.9 \exp(-y_{d1}^2(k-1)))y_{d1}(k-2) \\ y_{d2}(k) &= 0.6 y_{d2}(k-1) + 0.2 y_{d2}(k-1) y_{d2}(k-2) \\ &\quad + 1.2 \tanh(y_{d1}(k-2)) \end{aligned} \tag{39}$$

Given the initial conditions $y_{d1}(0) = y_{d1}(-1) = y_{d2}(0) = y_{d2}(-1) = 0.1$, the response of this noise-free time series is depicted in Fig. 1. A two-output RBF network was used to model this time series, with the input vector to the RBF network given by

$$\mathbf{x}(k) = [y_1(k-1) y_1(k-2) y_2(k-1) y_2(k-2)]^T \tag{40}$$

As each training input was used as a candidate RBF center, the number of candidate regressors in the RBF model (6) was $M = 500$.

For the multi-output modelling, the covariance of the modelling error $\mathbf{E}$, $\text{Cov}(\mathbf{E}) = \mathbf{E}^T\mathbf{E}$, is a $n_0 \times n_0$ matrix. Typical scalar measures of modelling accuracy include trace(Cov($\mathbf{E}$)) and det(Cov($\mathbf{E}$)). Since det(Cov($\mathbf{E}$)) is well-known to be a better measure of modelling accuracy, we will adopt the following scalar measure:

$$s_m = \log(\det(\text{Cov}(\mathbf{E}))) \tag{41}$$

in our modelling comparison. Table 1 compares the values of s_m over the training and testing sets for the RBF models constructed by the combined LROLS and D-optimality algorithm with those of the combined OLS and D-optimality algorithm, given a wide range of β values. For this example the true system noise $\epsilon(k)$ had a $s_m = -6.43775$. It can be seen that using the D-optimality alone without regularisation the constructed models can still fit into the noise unless the weighting β is set to some appropriate value. Combining regularisation with D-optimality design, the results obtained are consistent over a wide range of β values and, effectively, the value of β has no serious influence on the model construction process. The generalisation capability of an identified model can best be tested by examining the iterative model output. If the

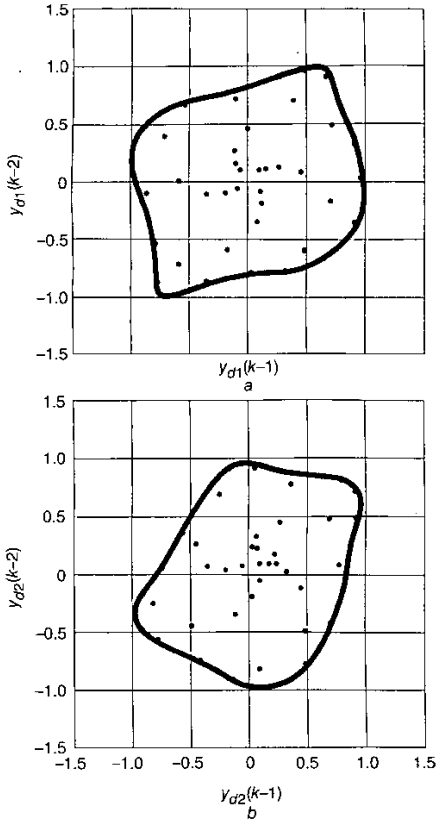

Fig. 1 *Two-dimensional representation of noise-free time series observations*

Initial conditions $y_{d1}(0)=y_{d1}(-1)=y_{d2}(0)=y_{d2}(-1)=0.1$
a Phase plot of noise free time series $y_{d1}(k)$
b Phase plot of noise free time series $y_{d2}(k)$

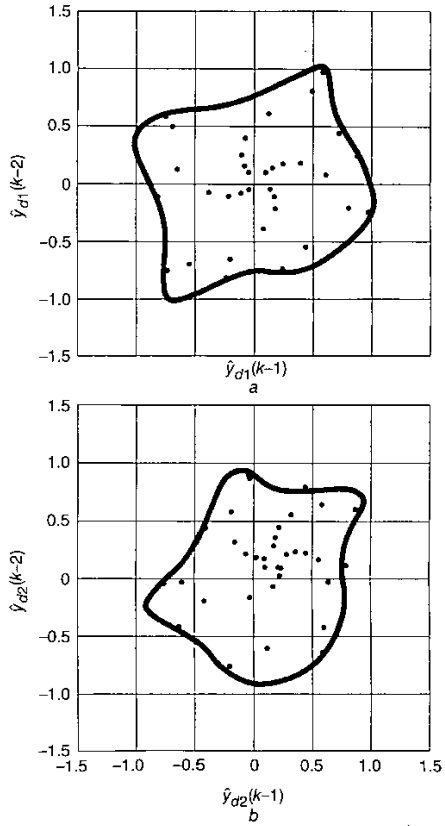

Fig. 2 *Two-dimensional representation of iterative model outputs*

Initial conditions $\hat{y}_{d1}(0)=\hat{y}_{d1}(-1)=\hat{y}_{d2}(0)=\hat{y}_{d2}(-1)=0.1$
49-term RBF model was constructed by combined LROLS and D-optimality algorithm with $\beta=1.0$ from very noisy data.
a Phase plot of iterative model output $\hat{y}_{d1}(k)$
b Phase plot of iterative model output $\hat{y}_{d2}(k)$

iterative model output can closely realise the behaviour shown in Fig. 1, the identified model truly captures the underlying dynamics of the system and does not simply fits the noise containing in the training data. Given the same initial conditions, the 49-term RBF model identified by the combined LROLS and D-optimality algorithm with $\beta=1.0$ were used to iteratively generate the network outputs $\hat{y}_{di}(k)$, $i=1, 2$, with the input

$$\mathbf{x}_d(k)=[\hat{y}_{d1}(k-1)\hat{y}_{d1}(k-2)\hat{y}_{d2}(k-1)\hat{y}_{d2}(k-2)]^T \quad (42)$$

The iterative model outputs so generated are plotted in Fig. 2. The constructed RBF model appeared to capture the underlying dynamics of the system well.

4.2 Simulated single-input two-output nonlinear system

The data were generated using the model

$$\begin{aligned} y_1(k) &= 0.5y_1(k-1)+u(k-1)+0.4\tanh(u(k-2)) \\ &\quad +0.1\sin(\pi y_1(k-2))y_2(k-1)+\epsilon_1(k) \\ y_2(k) &= 0.3y_2(k-1)+0.1y_2(k-2)y_1(k-1) \\ &\quad +0.4\exp(-u^2(k-1))y_1(k-2)+\epsilon_2(k) \end{aligned} \quad (43)$$

where the system input $u(k)$ was uniformly distributed in $(-0.5, 0.5)$, and the system noises $\epsilon(k)=[\epsilon_1(k)\ \epsilon_2(k)]^T$ were gaussian with zero means and covariance $0.04\mathbf{I}_2$. The data set contained 1000 samples, with the first 500 data points

Table 1: Comparison of modelling accuracy for simulated two-output nonlinear time-series modelling example. Cov(E): one-step prediction error covariance

D-optimality weighting β	Training set log(det(Cov(E)))		Testing set log(det(Cov(E)))		Number of terms	
	LROLS + D-opt	OLS + D-opt	LROLS + D-opt	OLS + D-opt	LROLS + D-opt	OLS + D-opt
0.001	−6.78104	−18.1385	−6.07734	−5.3000	102	470
0.01	−6.68156	−10.1001	−6.08521	−5.39079	62	302
0.1	−6.55440	−6.87149	−6.09854	−5.95289	50	72
1.0	−6.43524	−6.51637	−6.03528	−6.04794	49	49
10.0	−6.38538	−6.43935	−6.12874	−6.10428	44	44

Table 2: Comparison of modelling accuracy for simulated single-input two-output nonlinear system example. Cov(E): one-step prediction error covariance

D-optimality weighting β	Training set log(det(Cov(E)))		Testing set log(det(Cov(E)))		Number of terms	
	LROLS + D-opt	OLS + D-opt	LROLS + D-opt	OLS + D-opt	LROLS + D-opt	OLS + D-opt
0.01	−6.59701	−10.8873	−6.10548	−5.41334	44	320
0.1	−6.56962	−6.84887	−6.07789	−5.95589	38	61
1.0	−6.49324	−6.56252	−6.13198	−6.08903	35	36
10.0	−6.50340	−6.55698	−6.11586	−6.06297	35	35

Table 3: Comparison of modelling accuracy for simulated single-input two-output nonlinear system example. Cov($\mathbf{E}_d$): model iterative error covariance over entire 1000- sample data set

D-optimality weighting β	Log(det(Cov($\mathbf{E}_d$)))		Number of terms	
	LROLS + D-opt	OLS + D-opt	LROLS + D-opt	OLS + D-opt
0.01	−5.65089	−5.37460	44	320
0.1	−5.66776	−5.65160	38	61
1.0	−5.65614	−5.71936	35	36
10.0	−5.72100	−5.70334	35	35

used for training and the last 500 data samples for model validation. A two-output RBF network with the input

$$\mathbf{x}(k) = [y_1(k-1)y_1(k-2)y_2(k-1)y_2(k-2)u(k-1) \quad u(k-2)]^T \tag{44}$$

was employed to fit the noisy training data. The goodness of a fitted model was also evaluated by computing the iterative model outputs with the input

$$\mathbf{x}_d(k) = [\hat{y}_{d1}(k-1)\hat{y}_{d1}(k-2)\hat{y}_{d2}(k-1)\hat{y}_{d2}(k-2) \quad u(k-1)u(k-2)]^T \tag{45}$$

For this example, the true system noise again had $s_m = -6.43775$. The modelling accuracies over both the training and testing sets are compared in Table 2 for the two algorithms, the combined LROLS and D-optimality and the combined OLS and D-optimality, with a range of β values. Again it is seen that, for the combined LROLS and D-optimality algorithm, the model construction process is insensitive to the value of β. The modelling accuracies in terms of log(det(Cov($\mathbf{E}_d$))) for the two algorithms are compared in Table 3, where Cov($\mathbf{E}_d$) denotes the covariance of the iterative model error. The one-step predictions $\hat{\mathbf{y}}(k)$ of the 35-term RBF model produced by the combined

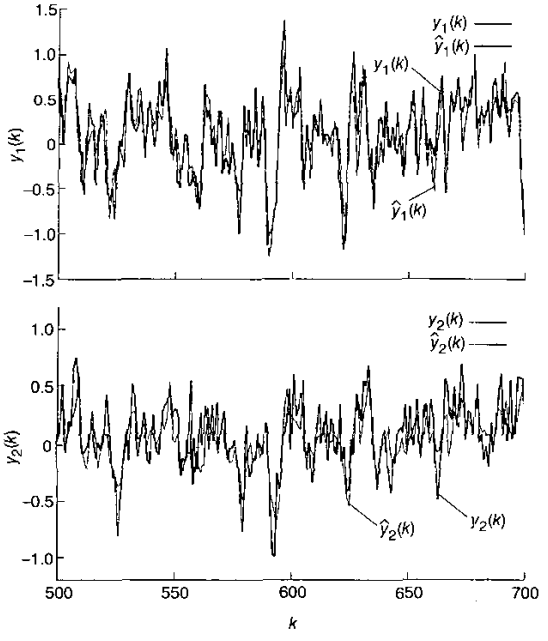

Fig. 3 *One-step prediction $\hat{y}(k)$ superposed on system output $y(k)$ over first 200 samples of test set for simulated single-input two-output nonlinear system*

35-term RBF model was identified by combined LROLS and D-optimality algorithm with $\beta = 10.0$

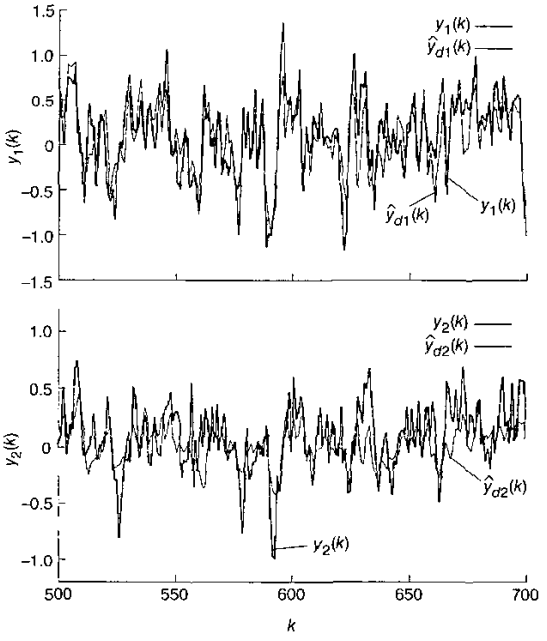

Fig. 4 *Model iterative output $\hat{y}_d(k)$ superposed on system output $\hat{y}(k)$ over first 200 samples of test set for simulated single-input two-output nonlinear example*

35-term RBF model was identified by combined LROLS and D-optimality algorithm with $\beta = 10.0$

Table 4: Comparison of modelling accuracy for turbo-alternator modelling example. Cov(E): one-step prediction error covariance, and Cov(E_d); model iterative error covariance

D-optimality weighting β	Training set log(det(Cov(E)))		Training set log(det(Cov(E_d)))		Number of terms	
	LROLS + D-opt	OLS + D-opt	LROLS + D-opt	OLS + D-opt	LROLS + D-opt	OLS + D-opt
0.00001	−18.4925	−28.2112	−13.3163	−27.6729	64	96
0.0001	−16.5032	−20.8628	−13.7963	−18.0451	49	78
0.001	−15.2006	−15.7269	−13.4131	−13.4300	34	40

LROLS and D-optimality algorithm with $\beta = 10.0$ are illustrated in Fig. 3, and the iterative model outputs $\hat{\mathbf{y}}_d(k)$ generated by the same RBF model are shown in Fig. 4.

4.3 Two-input two-output data set collected from turbo-alternator (appendix A11.3 in [26])

The data set contained 100 samples. The system inputs were the in-phase current deviation $u_1(k)$ and the out-of-phase current deviation $u_2(k)$, and the system outputs were the voltage deviation $y_1(k)$ and the frequency deviation $y_2(k)$. The two-output RBF network with the input vector

$$\mathbf{x}(k) = [y_1(k-1)y_1(k-2)y_1(k-3)y_2(k-1)y_2(k-2) \quad y_2(k-3)u_1(k-1)u_1(k-2)u_2(k-1) \quad u_2(k-2)]^T \tag{46}$$

was used to fit this data set. As the data set was too short to be divided into a training set and a testing set, the model validation in this case could only be performed by evaluating the iterative model outputs $\hat{y}_{di}(k)$, $i = 1, 2$, with the input

$$\mathbf{x}_d(k) = [\hat{y}_{d1}(k-1)\hat{y}_{d1}(k-2)\hat{y}_{d1}(k-3)\hat{y}_{d2}(k-1) \quad \hat{y}_{d2}(k-2)\hat{y}_{d2}(k-3)u_1(k-1)u_1(k-2) \quad u_2(k-1)u_2(k-2)]^T \tag{47}$$

over the training set of 100 samples. Table 4 compares the training accuracies of the two algorithms, the combined LROLS and D-optimality and the combined OLS and D-optimality, given three values of β. Although there were no statistics over a testing data set to confirm the generalisation capability of a resulting model, it can be seen from Table 4 that the combined LROLS and D-optimality algorithm performed more consistently with different β values. Note that with $\beta = 0.001$, the two algorithms had similar training accuracies, suggesting that the corresponding models should have similarly good generalisation capability. Figs. 5 and 6 depicted the model one-step predictions and the iterative model outputs, respectively, over the training data for the 34-term RBF model constructed by the combined LROLS and D-optimality algorithm with $\beta = 0.001$.

5 Conclusions

A locally regularised OLS algorithm with the D-optimality design has been proposed for constructing sparse multioutput RBF network models. The efficiency of the subset model selection procedure is ensured as usual with the orthogonal forward regression. By combining the two effective and complementary approaches for sparse and robust modelling, namely the local regularisation and

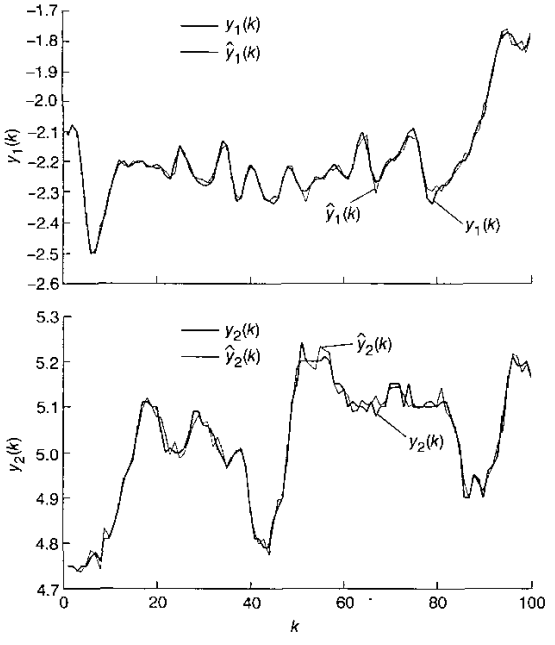

Fig. 5 *One-step prediction $\hat{y}(k)$ superposed on system output $y(k)$ for turbo-alternator modelling example*

34-term RBF model was identified by combined LROLS and D-optimality algorithm with $\beta = 0.001$

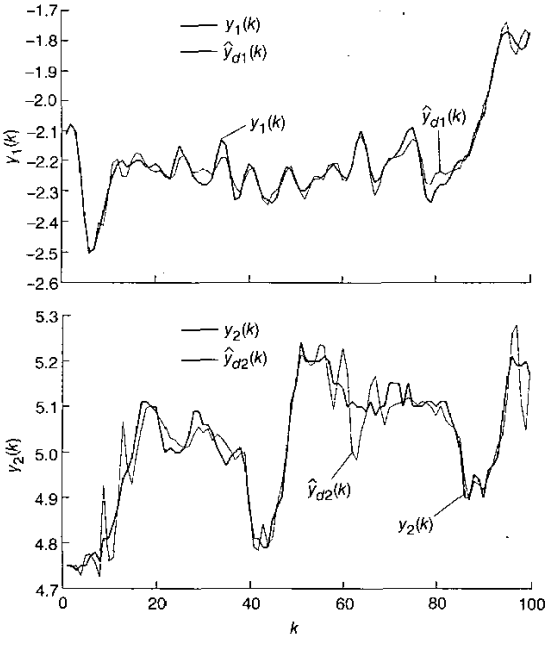

Fig. 6 *Model iterative output $\hat{y}_d(k)$ superposed on system output $y(k)$ for turbo-alternator modelling example*

34-term RBF model identified by combined LROLS and D-optimality algoritm with $\beta = 0.001$

D-optimality experimental design, the end result is an effective construction algorithm that is capable of producing sparse multioutput RBF network models with excellent generalisation performance. It has been shown that the performance of the algorithm is insensitive to the D-optimality cost weighting, and the model construction process is fully automated. The complexity of this combined model construction procedure is only slightly more than that of the efficient OLS algorithm.

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